

ON AN ASYMPTOTIC RELATIONSHIP BETWEEN SYMMETRIC BETA AND NORMAL DISTRIBUTIONS

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In this paper, which is a follow-up publication to [4], we show that normal distributions can be considered as weak limits of appropriately rescaled symmetric beta distributions.

Let X be a random variable with a symmetric standard beta distribution, i.e., with a p.d.f. f given by

$$f = (x; \alpha) = \frac{x^{\alpha-1} \cdot (1-x)^{\alpha-1}}{B(\alpha, \alpha)}, \quad 0 < x < 1; \quad \alpha > 0,$$

where $B(\alpha, \beta) = \frac{\Gamma(\alpha) \cdot \Gamma(\beta)}{\Gamma(\alpha + \beta)}$ denotes Euler's Beta function (cf. [2], (25.2)).

The first moments are given by

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$$E(X) = \frac{1}{2}, \quad \text{Var}(X) = \frac{1}{4 \cdot (2\alpha + 1)} \quad (\text{cf. [2], relations (25.15a) and (25.15b)}).$$

Now let $Y := c \cdot (2X - 1)$ with a real $c > 0$. Then Y follows a generalized symmetric beta distribution with support $[-c, c]$ and p.d.f.

$$g(y; \alpha, c) = \frac{1}{2c} f\left[\frac{1}{2} + \frac{y}{2c}, \alpha\right], \quad -c < y < c \quad (\text{cf. [2], (25.3) and (25.5)}).$$

For the first moments, we thus obtain

$$E(Y) = 0, \quad \text{Var}(Y) = \frac{c^2}{(2\alpha + 1)}.$$

For the sequel, we need the following Lemma.

Lemma. *There holds $k(\alpha) < B(\alpha, \alpha) < k(\alpha) \cdot e^{1/(6\alpha)}$ with $k(\alpha) = 2\sqrt{\frac{\pi}{\alpha}} \cdot 4^{-\alpha}$, $\alpha > 0$ and hence $0 < B(\alpha, \alpha) - k(\alpha) < \frac{k(\alpha)}{6\alpha} \cdot \exp\left(\frac{1}{6\alpha}\right)$ with $\lim_{\alpha \rightarrow \infty} \frac{B(\alpha, \alpha)}{k(\alpha)} = 1$.*

Proof. By the known analytical estimate for Stirling's formula (cf. [1], (6.1.38))

$$\sqrt{\frac{2\pi}{x}} \left(\frac{x}{e}\right)^x < \Gamma(x) < \sqrt{\frac{2\pi}{x}} \left(\frac{x}{e}\right)^x \cdot \exp\left(\frac{1}{12x}\right), \quad x > 0$$

there follows

$$\frac{2\pi}{x} \left(\frac{x}{e}\right)^{2x} < \Gamma^2(x) < \frac{2\pi}{x} \left(\frac{x}{e}\right)^{2x} \cdot \exp\left(\frac{1}{6x}\right)$$

$$\text{and } \sqrt{\frac{\pi}{x}} \left(\frac{2x}{e}\right)^{2x} < \Gamma(2x) < \sqrt{\frac{\pi}{x}} \left(\frac{2x}{e}\right)^{2x} \cdot \exp\left(\frac{1}{24x}\right),$$

hence

$$0 < \exp\left(-\frac{1}{24\alpha}\right) < \frac{B(\alpha, \alpha)}{k(\alpha)} = \frac{\Gamma^2(\alpha)}{\Gamma(2\alpha) \cdot k(\alpha)} < \exp\left(\frac{1}{6\alpha}\right), \quad \alpha > 0$$

from which the Lemma follows immediately. Note that $\exp(x) - 1 < x \cdot \exp(x)$, $x > 0$.

The following figure shows the function $h(\alpha) = \frac{B(\alpha, \alpha) - k(\alpha)}{k(\alpha)} \cdot 6(\alpha)$,

$0 < \alpha < 100$:

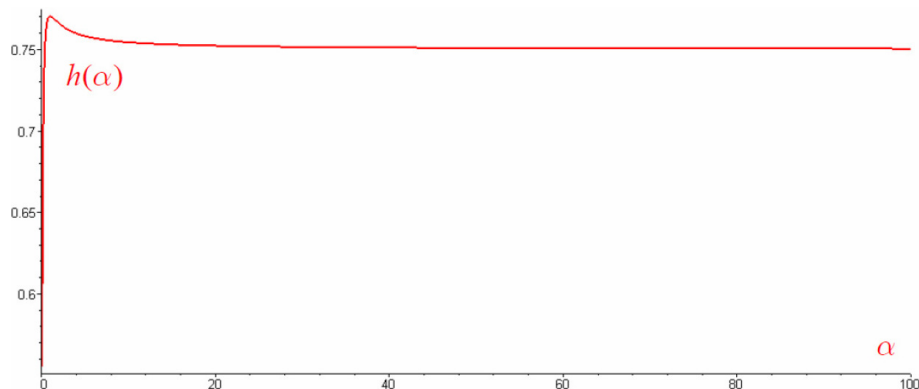


Figure 1. Plot of $g(\alpha)$.

This leads to the following

Theorem. Let $\sigma > 0$ be fixed and $c := \sqrt{2\alpha} \cdot \sigma$ for $\alpha > 0$. Then

$$\lim_{\alpha \rightarrow \infty} g(y; \alpha, c) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{y^2}{2\sigma^2}\right), \quad \text{i.e., the rescaled symmetric beta}$$

distribution converges weakly to the normal distribution with mean zero and variance σ^2 .

Proof. We have

$$g(y; \alpha, c) = \frac{1}{2c} f\left(\frac{1}{2} + \frac{y}{2c}, \alpha\right) = \frac{\left(\frac{1}{2} + \frac{y}{2c}\right)^{\alpha-1} \cdot \left(\frac{1}{2} - \frac{y}{2c}\right)^{\alpha-1}}{B(\alpha, \alpha) \cdot 2c}$$

$$\begin{aligned}
&= \frac{\left(1 - \frac{y^2}{c^2}\right)^{\alpha-1}}{4^{\alpha-1} \cdot k(\alpha) \cdot 2c} \cdot \frac{k(\alpha)}{B(\alpha, \alpha)} = \dots \\
&\dots = \frac{\left(1 - \frac{y^2}{c^2}\right)^{\alpha-1}}{\sqrt{2\pi\sigma}} \cdot \frac{k(\alpha)}{B(\alpha, \alpha)}, \quad -c < y < c.
\end{aligned}$$

Now since $\lim_{\alpha \rightarrow \infty} \left(1 - \frac{y^2}{c^2}\right)^{\alpha-1} = \lim_{\alpha \rightarrow \infty} \left(1 - \frac{y^2}{2\alpha\sigma^2}\right)^{\alpha-1} = \exp\left(-\frac{y^2}{2\sigma^2}\right)$, the

Theorem is proved.

The following figures show some comparisons between $g(y; \alpha, c)$ (plotted in red) and $n(y; \sigma) = \frac{1}{\sqrt{2\pi\sigma}} \exp\left(-\frac{y^2}{2\sigma^2}\right)$ (plotted in blue), for several choices of α and σ .

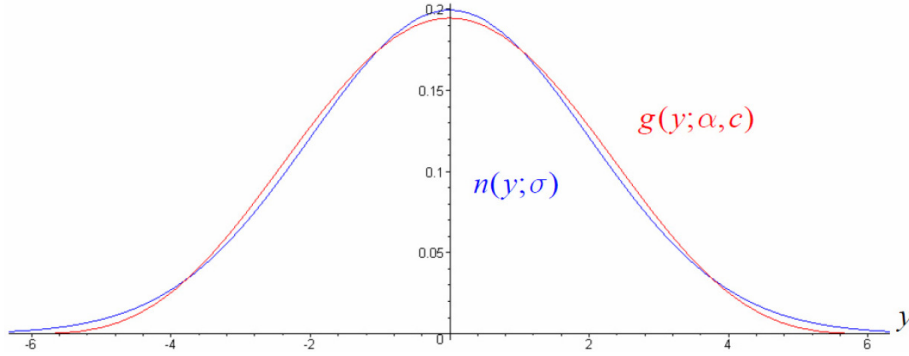


Figure 2. $\alpha = 10$, $\sigma = 2$, $c = 8.944\dots$

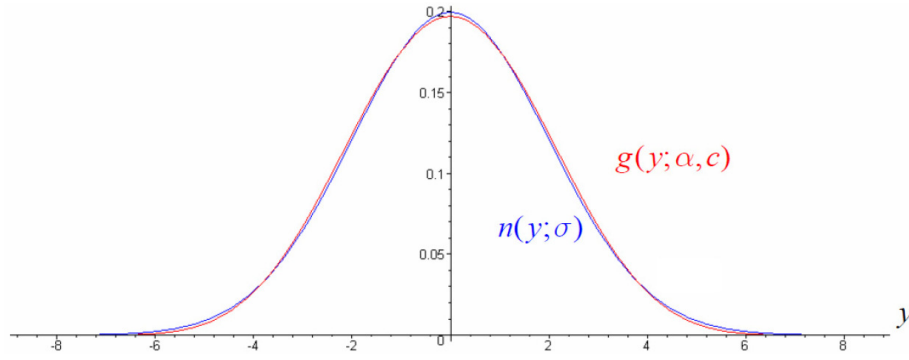


Figure 3. $\alpha = 5$, $\sigma = 2$, $c = 6.324\dots$.

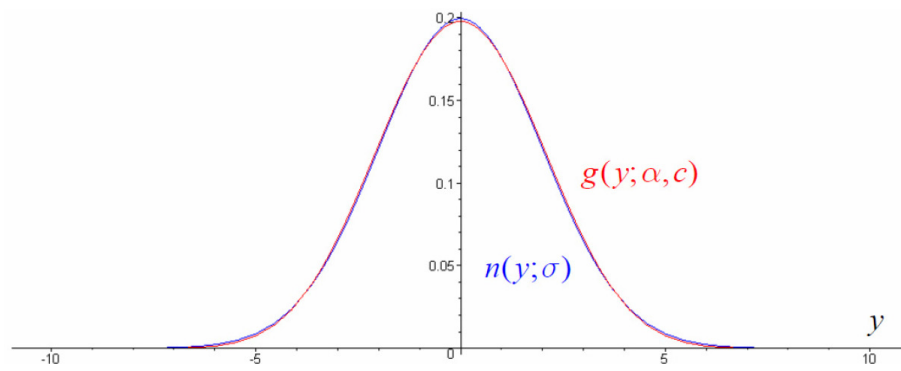


Figure 4. $\alpha = 15$, $\sigma = 2$, $c = 10.954\dots$.

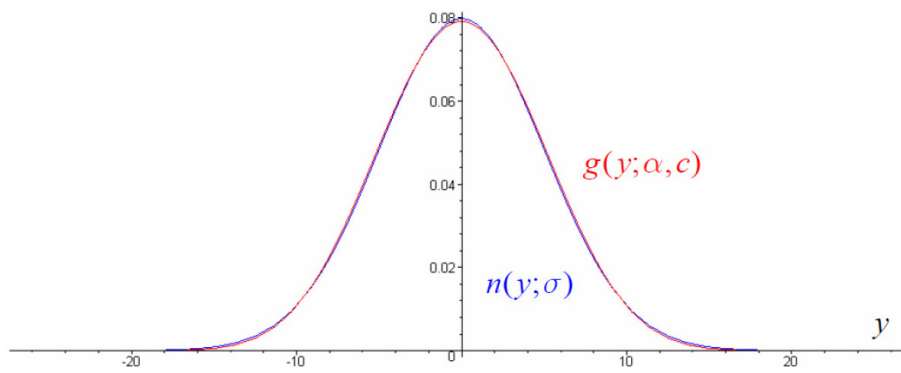


Figure 5. $\alpha = 15$, $\sigma = 5$, $c = 27.386\dots$.

Note that for the variance of Y , we have $Var(Y) = \frac{c^2}{(2\alpha + 1)}$
 $= \frac{2\alpha\sigma^2}{(2\alpha + 1)} \approx \sigma^2$ for large values of α .

Further relationships between beta and other distributions are discussed in [3].

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